

Incremental Healthcare Utilization and Cost Burden of Comorbid Insomnia in Alzheimer's Disease Patients

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Abstract.

Background: Insomnia is associated with worsened clinical outcomes among Alzheimer's disease dementia (AD) patients, increased caregiver burden, and healthcare utilization.

Objective: This study aimed to characterize the incremental healthcare burden of insomnia in AD using real-world data.

Methods: A retrospective observational study was conducted on AD patients selected from the IBM® MarketScan Commercial and Medicare Supplemental Databases. AD patients with claims-based evidence of insomnia were direct matched to a non-insomnia cohort based on demographic factors. Healthcare utilization and associated costs were assessed for a 12-month follow-up period.

Results: A total of 3,500 insomnia AD patients and 9,884 non-insomnia AD patients were analyzed. The insomnia cohort had a higher comorbidity burden at baseline (mean score on Charlson Comorbidity Index 2.5 versus 2.2, $p < 0.001$) and higher proportions of patients with baseline diagnoses for other conditions including depression: 40%, insomnia cohort versus 25%, non-insomnia ($p < 0.001$). AD patients with insomnia were more likely to have a claim for inpatient hospitalizations (39.8% versus 32.3%), emergency room services (56.4% versus 48.0%), and skilled-nursing services (42.6% versus 31.9%) (all $p < 0.05$). Mean total annual healthcare costs during the 12-month follow-up period were significantly higher among AD patients with insomnia as compared to those without. (Mean costs: \$37,356 versus \$27,990, $p < 0.001$).

Conclusion: AD patients with comorbid insomnia are more likely to use higher-cost healthcare services such as inpatient hospitalization, and skilled nursing, and have higher total healthcare costs. This real-world analysis provides evidence that AD disease management should consider proper treatment of comorbid insomnia due to the incremental burden and cost implications.

Keywords: Alzheimer's disease, dementia, facilities and services utilization, health care economics, sleep initiation and maintenance disorders

INTRODUCTION

Alzheimer's disease (AD) is a neurodegenerative disease and the leading cause of dementia in the elderly [1, 2]. The Alzheimer's Association estimated there were 5.8 million Americans age 65 and older living with AD in 2019 [3], 81% of which were at

least 75 years old [4]. The population of people over 65 in the United States living with AD is projected to grow to 12.7–13.8 million by 2050 [4, 5].

The classic hallmarks of AD are progressive deterioration of memory, intellect, and motor skills leading to increasing disability and dependency on caregivers [3, 6]. Patients with AD have increased use of healthcare services, including more frequent high-cost events like hospitalizations and emergency room visits, compared to cognitively normal individuals [7]. This increased need for care creates a significant

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financial burden. In 2019 the costs associated with healthcare and long-term care for individuals with dementia due to AD and other causes were estimated to be \$290 billion, with the average Medicare beneficiary with AD costing \$35,000 more annually than Medicare beneficiaries without AD [3].

Sleep disorders such as insomnia occur more frequently in AD patients than in the general population [8]. Between 25%–40% of patients with AD report having some type of sleep disorder, such as excessive awakenings, early morning awakenings, excessive daytime sleepiness, and napping for more than 1 h during the day [9, 10]. A recent meta-analysis found that 20% of patients with dementia living in hospice care had clinically significant sleep disturbances, 38% had any symptom of sleep disturbance, and 70% had poor sleep efficiency as measured by actigraphy [11].

The relationship between sleep and AD is theorized to be bidirectional, with poor sleep as both a cause and consequence of AD [12]. Sleep disturbances often appear early in the course of AD progression and tend to be correlated with the rate of cognitive decline [13]. Conversely, poor sleep in older adults may result in cognitive decline that could progress to AD [14, 15].

In patients with AD, sleep disturbances are associated with worsening behavioral and psychological symptoms, which may lead to increased caregiver burden, increased fall risk, and increased use of high-cost interventions such as skilled nursing support and early institutionalization [16, 17]. However, no studies have looked at the healthcare resource utilization and cost implications of insomnia in AD. In this study, we characterized the incremental healthcare burden of insomnia in AD using real-world data by comparing all-cause healthcare utilization and costs between a cohort of AD patients with insomnia to a matched cohort of AD patients without insomnia. Multivariable analysis was used to estimate all-cause total healthcare costs, odds of an inpatient admission, and odds of having a claim for skilled nursing facility services after adjusting for baseline differences between AD patients with and without insomnia.

METHODS

Study design and data source

This retrospective cohort study utilized de-identified US administrative claims data from the IBM MarketScan Commercial Claims and Encounters

database (Commercial) and MarketScan Medicare Supplemental and Coordination of Benefits database (Medicare) for the period from January 1, 2014 to December 31, 2018. The Commercial database contains the inpatient, outpatient, and outpatient prescription drug experience of employees and their dependents, covered under a variety of fee-for-service and managed care health plans. The Medicare database contains the same healthcare experience of retirees with Medicare supplemental insurance paid for by employers. Both the Medicare-covered portion of payment (represented as Coordination of Benefits Amount) and the employer-paid portion are included in this database.

All database records are statistically de-identified and certified to be fully compliant with US patient confidentiality requirements set forth in the Health Insurance Portability and Accountability Act of 1996. Because this study used only de-identified patient records and did not involve the collection, use, or transmittal of individually identifiable data, this study was exempted from Institutional Review Board approval.

Relevant diagnoses, procedures, and treatments were obtained using International Classification of Diseases, 9th and 10th Revision, Clinical Modification (ICD-9-CM and ICD-10-CM) codes, Current Procedural Terminology 4th edition codes, Healthcare Common Procedure Coding System codes, and National Drug Codes.

Patient selection and cohort assignment

We identified patients with at least one inpatient or two outpatient claims with a diagnosis of AD (ICD-9-CM: 331.0; ICD-10-CM: G30.0, G30.1, G30.8, G30.9) between January 1, 2015 and December 31, 2018 in the MarketScan databases. The date of the first claim with a diagnosis of AD was defined as the AD diagnosis date. To be eligible for study inclusion, we required that patients were 50–100 years old on the AD diagnosis date, had continuous enrollment with medical and pharmacy benefits for at least 12 months before and after the AD diagnosis date, and had no claims with a diagnosis of AD in the 12 months before the AD diagnosis date.

From this population of patients newly-diagnosed with AD, we identified two populations: those with a diagnosis of insomnia or claim for an insomnia medication in the three months preceding or 12 months following the AD diagnosis date (insomnia cohort) and those without a diagnosis of insomnia

or claim for an insomnia medication in the same period (non-insomnia cohort). The diagnosis codes for insomnia used in this analysis included ICD-9-CM codes 327.00, 327.01, 327.02, 780.51, and 780.52 along with ICD-10-CM codes F51.01, F51.02, F51.03, F51.04, F51.05, F51.09, F51.8, F51.9, G47.00, G47.01, and G47.09. Prescription insomnia medications included benzodiazepines (estazolam, flurazepam, quazepam, temazepam, and triazolam) and non-benzodiazepines (low-dose doxepin [3–6 mg], eszopiclone, ramelteon, suvorexant, zaleplon, and zolpidem tartrate). The index date for the insomnia cohort was the date of the first claim with an insomnia diagnosis for all patients with an insomnia diagnosis or the date of the first claims for an insomnia medication for patients without an insomnia diagnosis but with a claim for prescription insomnia medication. The index date of controls was assigned after matching and set as the date corresponding to the same number of days between the matched insomnia patient's AD diagnosis date and insomnia diagnosis date.

We required that patients have continuous enrollment with medical and pharmacy benefits for at least 12 months before and after the index date, and we excluded patients with a diagnosis of a non-insomnia sleep disorder in the three months preceding or 12 months following the AD diagnosis date. Where possible, insomnia patients were directly matched (1:3) to non-insomnia patients on age, sex, region, and year of the index insomnia diagnosis. Due to sample limitations, it was not possible to identify three non-insomnia matched controls with sufficient continuous enrollment for every insomnia patient. Insomnia patients with no matches were excluded, while those with 1 or 2 matches were retained in the sample set.

Outcomes

Demographic characteristics were measured on the index date and included age and sex. Baseline clinical characteristics were measured during the 12-months preceding the index date and included Deyo-Charlson Comorbidity Index (DCI) and the presence of selected comorbid conditions representing common cost drivers and common neurological and psychological conditions, including cerebrovascular disease, congestive heart failure, depression, epilepsy, hyperlipidemia, hypertension, malignancy, mild liver disease, moderate or severe liver disease, mood disorder, myocardial infarction, other

sleep disorders, Parkinson's disease, peripheral vascular disease, pneumonia, renal disease, sleep apnea, and substance abuse disorders. All-cause healthcare resource utilization, including inpatient admissions, outpatient services (emergency department [ED] visits, physician office visits, skilled nursing facility services, hospice services, and other outpatient services), and outpatient pharmacy services were recorded in the 12 months following the index date. All-cause costs for these claims were calculated from the paid amounts of fully adjudicated claims, including insurer and health plan payments, as well as patient cost-sharing in the form of copayment, deductible, and coinsurance. All costs are reported as per person per month (PPPM) and were inflated to 2018 US dollars using the medical care component of the Consumer Price Index [18].

Statistical analysis

Mean and standard deviation (SD) were reported for continuous variables, and statistical significance was determined using Student's *t*-tests. Frequencies and percentages were reported for categorical variables, and statistical significance was determined using Chi-square tests. The alpha level for all statistical tests was 0.05. All data analyses were conducted using WPS version 4.2 (World Programming, United Kingdom).

Multivariable analysis was used to examine primary healthcare utilization and cost outcomes while controlling for baseline demographic and clinical characteristics. The outcomes analyzed in separate generalized linear models included total all-cause healthcare costs, odds of an inpatient admission, and costs of a claim for skilled nursing facility services. Logistic regression was used to model the odds of an inpatient admission or skilled nursing facility claim occurring in the follow-up period. Costs were analyzed using a generalized linear model (GLM) with a log link function and a gamma error distribution. The main explanatory variable of interest was the presence of insomnia diagnosis. Each model adjusted for age, sex, and all baseline comorbid conditions listed in the outcomes section except epilepsy (due to rarity and similarity across cohorts), malignancy and renal disease (both due to extreme rarity in the cohorts).

To estimate the costs from the model in terms of dollars, two approaches were taken. The first was to use the dollar costs predicted by the model. The second was to use a counterfactual based approach where two copies of the full data set were made, with one

being labelled the AD control group and the other the insomnia group. For the AD control copy the cost for those people who were in the insomnia group was changed to a counterfactual estimate of AD control cost based on the GLM coefficient only; and vice versa with the insomnia copy, i.e., the cost for the AD control people was changed to the counterfactual estimate for the insomnia group. Summary statistics for each of these copies were then used for estimates of the dollar costs for each person if they had had been in the other group. In most cases the counterfactual estimates of the group means were only modestly different from those observed or estimated from the model predictions.

RESULTS

There were 116,541 individuals in the MarketScan databases with a qualifying diagnosis of AD between January 1, 2015 and December 31, 2018 (Fig. 1). Of these, 112,466 were between the ages of 50 and 99 at their first AD diagnosis during the selection window. We excluded 61,128 individuals because they did not meet the initial continuous enrollment criteria and an additional 14,473 individuals because they had a prior diagnosis of AD. This left 36,865 qualifying adults with AD, of which 5,374 (14.6%) had a diagnosis of insomnia or claim for an insomnia medication in the 3 months preceding or 12 months following their AD diagnosis.

After excluding patients with evidence of other sleep disorders and those who did not have 12 months of continuous enrollment in MarketScan before and after their qualifying insomnia diagnosis or assigned proxy diagnosis date, there were 3,602 AD patients with insomnia and 25,312 AD patients without insomnia. The patients with insomnia were then directly matched on age, sex, region, and year of diagnosis to up to three patients without insomnia. AD patients with insomnia not directly matched to at least one non-insomnia patient were excluded from further analysis. This approach was used to balance patient characteristics and to maintain sample size of the insomnia cohort. A total of 3,500 AD patients with insomnia met all inclusion criteria and were matched to 9,884 AD patients without insomnia (Fig. 1). This reflects the loss of 102 AD patients due to lack of at least one direct match in the initial sample and an effective matching ratio of 1:2.82 among the remaining patients; Supplementary Table 1 shows the demographic characteristics of the pre-match group.

The mean age of AD patients in both the insomnia and non-insomnia cohorts was approximately 82 years, and about 67% of patients in both cohorts were female (Table 1). Patients with insomnia had a higher CCI in the baseline period compared to non-insomnia patients (2.50 versus 2.16, $p < 0.001$) and were more likely to have a prior diagnosis of many comorbid conditions of interest, most notably depression (39.5% versus 24.9%, $p < 0.001$), moderate to severe liver disease (36.8% versus 31.9%, $p < 0.001$), and hypertension (78.9% versus 74.3%, $p < 0.001$). Patients with insomnia were also more likely to have a prior diagnosis of hyperlipidemia (53.7% versus 50.9%, $p = 0.004$), peripheral vascular disease (25.7% versus 22.6%, $p < 0.001$), myocardial infarction (18.2% versus 15.5%, $p < 0.001$), mild liver disease (15.6% versus 12.2%, $p < 0.001$), congestive heart failure (13.9% versus 11.2%, $p < 0.001$), pneumonia (12.3% versus 8.5%, $p < 0.001$), Parkinson's disease (7.0% versus 5.2%, $p < 0.001$), substance abuse disorders (4.0% versus 1.9%, $p < 0.001$), mood disorders (4.0% versus 1.9%, $p < 0.001$), other sleep disorders (3.8% versus 2.1%, $p < 0.001$), or cerebrovascular disease (3.8% versus 2.9%, $p = 0.005$).

AD patients with insomnia had higher all-cause healthcare utilization than AD patients without insomnia (Table 2). Patients with insomnia were more likely to have at least one inpatient admission (39.8% versus 32.3%, $p < 0.001$), had more frequent inpatient admissions (mean [SD]: 0.6 [0.9] versus 0.4 [0.7], $p < 0.001$), and more days spent in the inpatient setting (9.1[13.5] versus 7.6[11.0], $p < 0.001$). Patients with insomnia were also more likely to have greater utilization of ED services (56.4% versus 48.0%, $p < 0.001$), skilled nursing facilities (42.6% versus 31.9%, $p < 0.001$), or other outpatient services (98.5% versus 97%, $p < 0.001$). While the percentage of patients with at least one outpatient office visit or hospice service claim was not significantly different between patients with or without insomnia, patients with insomnia had, on average, more office visits (mean [SD]: 8.8 [9.1] versus 7.5 [7.6], $p < 0.001$) and more hospice claims (0.3 [6.6] versus 0.1 [2.7], $p = 0.012$) than patients without insomnia. In addition, patients with insomnia were more likely to have at least one outpatient pharmacy claim (97.4% versus 95.3%, $p < 0.001$) and had, on average, more outpatient pharmacy claims (65.3 claims versus 47.8 claims, $p < 0.001$).

AD patients with insomnia had higher mean (SD) and median unadjusted all-cause healthcare costs PPPY than AD patients without insomnia

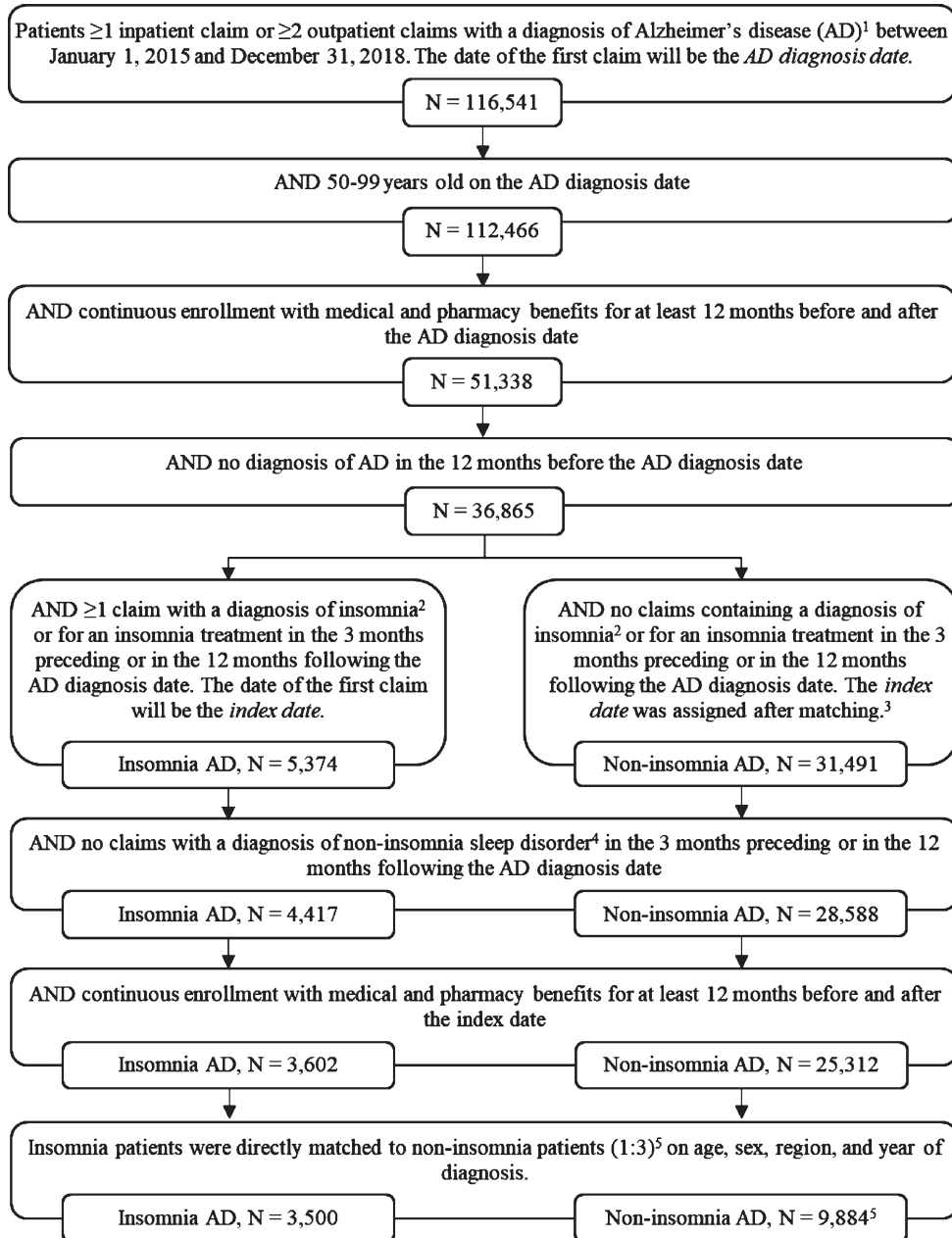


Fig. 1. Patient selection. ¹Alzheimer's disease diagnosis codes: ICD-9-CM code 331.0; ICD-10-CM codes: G30.0, G30.1, G30.8, G30.9. ²Insomnia diagnosis codes: ICD-9-CM code 331.0; ICD-10-CM codes: G30.0, G30.1, G30.8, G30.9. ³Control patients will be assigned an index date corresponding to the same number of days between the matched insomnia patient's AD diagnosis and first treatment for insomnia (e.g., -3 months to +12 months). ⁴Non-insomnia sleep disorders include sleep apnea, narcolepsy, and others. ⁵Due to sample limitations it was not possible to identify three non-insomnia matched controls with sufficient continuous enrollment for every insomnia patient. Insomnia patients with no matches were excluded while those with 1 or 2 matches were retained in the sample set.

(mean [SD]: \$37,356 [\$70,089] versus \$27,990 [\$57,179], $p < 0.001$); median costs: \$18,813 versus \$12,318) (Fig. 2 and Table 3). The largest driver of higher costs among patients with insomnia were inpatient claims (\$13,817 [\$48,854] versus \$9,657

[\$32,563], $p < 0.001$). Costs were also higher for outpatient services overall (\$18,546 [\$36,351] versus \$14,323 [\$4,432], $p < 0.001$), as well as all specific sub-categories of outpatient care except hospice: emergency room services (\$2,514 [\$8,932] versus

Table 1
Demographics and baseline clinical characteristics

	Insomnia patients		Non-insomnia patients		<i>p</i>
	<i>N</i> = 3,500		<i>N</i> = 9,884		
Age (Mean, SD)	81.7	8.6	82.0	8.2	0.077
Female	2,334	66.7%	6,651	67.3%	0.513
Calendar year of AD diagnosis date (<i>N</i> , %)					
2015	1,777	50.8%	5,057	51.2%	0.928
2016	1,223	34.9%	3,454	34.9%	
2017	499	14.3%	1,369	13.9%	
2018	1	0.0%	4	0.0%	
Charlson Comorbidity Index (Mean, SD)	2.50	2.23	2.16	2.10	<0.001
Comorbid conditions (<i>N</i> , %)					
Hypertension	2,761	78.9%	7,345	74.3%	<0.001
Hyperlipidemia	1,881	53.7%	5,035	50.9%	0.004
Depression	1,384	39.5%	2,460	24.9%	<0.001
Moderate or severe liver disease	1,287	36.8%	3,150	31.9%	<0.001
Peripheral vascular disease	899	25.7%	2,233	22.6%	<0.001
Myocardial infarction	636	18.2%	1,536	15.5%	<0.001
Mild liver disease	546	15.6%	1,208	12.2%	<0.001
Congestive heart failure	488	13.9%	1,106	11.2%	<0.001
Pneumonia	430	12.3%	844	8.5%	<0.001
Parkinson's disease	244	7.0%	514	5.2%	<0.001
Substance abuse disorders	141	4.0%	192	1.9%	<0.001
Mood disorder	140	4.0%	187	1.9%	<0.001
Other sleep disorders ¹	134	3.8%	208	2.1%	<0.001
Cerebrovascular disease	134	3.8%	283	2.9%	0.005
Epilepsy	132	3.8%	369	3.7%	0.919
Sleep apnea ¹	61	1.7%	135	1.4%	0.111
Any malignancy	27	0.8%	60	0.6%	0.298
Renal disease	12	0.3%	42	0.4%	0.510

Demographic information captured on index date. Clinical characteristics were measured during the 12-month pre-index baseline period.

¹This captures the subset of patients with a diagnosis of other sleep disorders or sleep apnea during the variable period between the start of the 12-month pre-index period and the start of the 3-month pre-AD diagnosis exclusion period.

\$1,730 [\$6,334], $p < 0.001$), outpatient office visits (\$988 [\$1,151] versus \$839 [\$1,007], $p < 0.001$), and skilled nursing facility claims (\$5,333 versus \$3,674, $p < 0.001$). The full cost results, including medians, can be found in Table 3.

After adjusting for baseline differences in baseline demographic and clinical characteristics, patients with insomnia had significantly higher odds (95%CI) of an inpatient admission (1.38 [1.22–1.44], $p < 0.001$) and of a skilled nursing facility claim (1.44 [1.32–1.56], $p < 0.001$) than patients without insomnia (Fig. 3). Similarly, patients with insomnia experienced 28% higher all-cause healthcare costs than patients without insomnia (1.28 [1.23–1.34], $p < 0.001$). Mean (95%CI) model-adjusted predicted costs were \$37,849 (\$37,514–\$38,183) for patients with insomnia and \$27,810 (\$27,670–\$27,949) for patients without insomnia. The full results of the multivariable models can be found in Supplementary Table 2.

DISCUSSION

While the research literature on AD and on sleep disorders are both extensive, relatively little work has explored the intersection of these patient populations, particularly regarding research on patterns of healthcare utilization and expenditures due to insomnia within the AD population. This research helps to fundamentally characterize this intersecting patient population, which may face unique clinical challenges and benefit from specific treatment strategies. They may also be a chronically underserved group, as insomnia is commonly underdiagnosed and undertreated among elderly populations [19].

In this analysis of the MarketScan databases, we found that 14.6% of patients newly diagnosed with AD had a diagnosis of insomnia or received a prescription treatment for insomnia in the 3 months preceding or 12 months following their AD diagnosis. This is consistent with prior claims analyses, which

Table 2
All-cause healthcare utilization in the 12-month follow-up period

	Insomnia patients		Non-insomnia patients		p
	N = 3,500		N = 9,884		
Inpatient					
Patients with ≥ 1 inpatient admission (N, %)	1,394	39.8%	3,194	32.3%	<0.001
Number of inpatient admissions ¹ (Mean, SD)	0.6	0.9	0.4	0.7	<0.001
Total days in inpatient setting (Mean, SD)	9.1	13.5	7.6	11.0	<0.001
Outpatient (OP)					
Emergency department (ED) visits					
Patients with ≥ 1 ED visit (N, %)	1,974	56.4%	4,744	48.0%	<0.001
Number of ED visits ¹ (Mean, SD)	1.5	2.3	1.0	1.6	<0.001
OP office visits					
Patients with ≥ 1 office visit (N, %)	2,888	82.5%	8,200	83.0%	0.546
Number of office visits ¹ (Mean, SD)	8.8	9.1	7.5	7.6	<0.001
Skilled nursing facility (SNF) services					
Patients with ≥ 1 SNF claim (N, %)	1,490	42.6%	3,154	31.9%	<0.001
Number of SNF claims ¹ (Mean, SD)	17.0	49.5	9.7	33.6	<0.001
Hospice services					
Patients with a hospice claim (N, %)	29	0.8%	61	0.6%	0.188
Number of hospice claims ¹ (Mean, SD)	0.3	6.6	0.1	2.7	0.012
Other OP services					
Patients with ≥ 1 other OP service (N, %)	3,448	98.5%	9,585	97.0%	<0.001
Number of other OP services ¹ (Mean, SD)	62.1	73.1	48.6	64.6	<0.001
OP pharmacy					
Patients with ≥ 1 OP claim prescription (N, %)	3,409	97.4%	9,417	95.3%	<0.001
Number of claims ¹ (Mean, SD)	5.3	45.2	47.8	40.5	<0.001

¹ Among all patients.

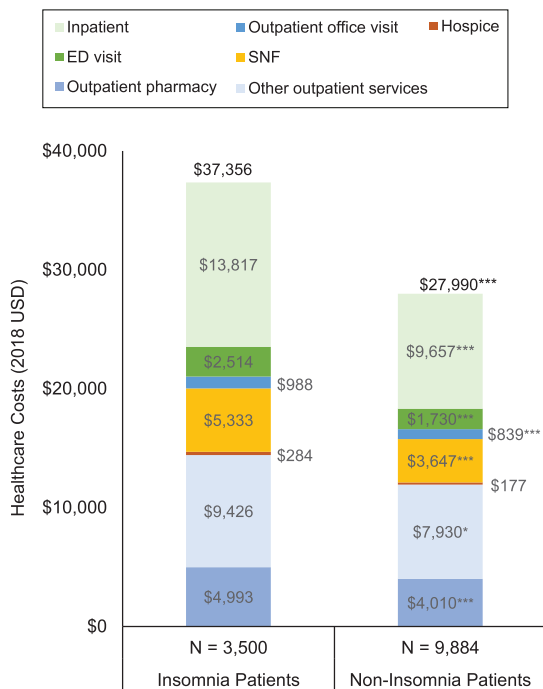


Fig. 2. All-cause healthcare costs per-person per-year in the 12-month follow-up period. ED, emergency department; SNF, skilled nursing facility. **p* < 0.05, ****p* < 0.001.

have reported that 12%–13% of patients with AD have a comorbid diagnosis of insomnia and 8%–17% have a recent prescription for an insomnia treatment [20–23]. Higher estimates of the prevalence of disturbed sleep among patients with AD have been found through caregiver interviews and other informant-based tools [9–11].

Patients included in this study were on average 82 years old, which is similar to other claims-based studies of newly-diagnosed AD [20–22]. However, a higher percentage were female (67% versus 53%–63%) [21–24], which may reflect the higher prevalence of insomnia among women, as found in other studies [8, 25]. In addition to being more common, some research suggests that in women insomnia presents in more complex ways with more symptoms [26] and that the adverse effects of insomnia and sleep loss are more severe among women compared to men with regards to cardiovascular health [27, 28], risk of diabetes [28], and psychological affect, functioning, and quality of life [28, 29].

Consistent with prior studies of comorbidity burden among adults with insomnia [8, 15], we found that AD patients with comorbid insomnia had higher rates of many comorbidities, including hypertension, hyperlipidemia, and depression, than AD patients

Table 3
All-cause healthcare costs in the 12-month follow-up period

	Insomnia patients		Non-insomnia patients		<i>p</i>
	<i>N</i> = 3,500		<i>N</i> = 9,884		
Total costs (Mean, SD)	\$37,356	\$70,089	\$27,990	\$57,179	<0.001
Median	\$18,813		\$12,318		
Inpatient costs (Mean, SD)	\$13,817	\$48,854	\$9,657	\$32,562	<0.001
Median	\$0		\$0		
Outpatient costs (Mean, SD)	\$18,546	\$36,351	\$14,323	\$39,928	<0.001
Median	\$8,633		\$5,432		
ED visit costs (Mean, SD)	\$2,514	\$8,932	\$1,730	\$6,334	<0.001
Median	\$192		\$0		
Outpatient office visit costs (Mean, SD)	\$988	\$1,151	\$839	\$1,007	<0.001
Median	\$666		\$588		
SNF costs (Mean, SD)	\$5,333	\$12,988	\$3,647	\$11,089	<0.001
Median	\$0		\$0		
Hospice costs (Mean, SD)	\$284	\$4,121	\$177	\$3,144	0.112
Median	\$0		\$0		
Other outpatient services costs (Mean, SD)	\$9,426	\$30,733	\$7,930	\$36,565	0.030
Median	\$3,873		\$2,789		
Outpatient pharmacy costs (Mean, SD)	\$4,993	\$8,220	\$4,010	\$6,876	<0.001
Median	\$2,981		\$2,216		

ED, emergency department; SNF, skilled nursing facility.

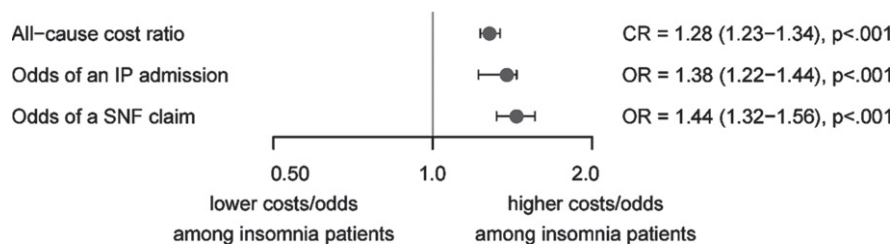


Fig. 3. Multivariable models of total healthcare costs, odds of inpatient (IP) admission, and odds of having a claim for skilled nursing facility (SNF) services with presence of insomnia diagnosis as the primary explanatory variable. CR, cost ratio; OR, odds ratio.

without insomnia. Overall, the demographic and clinical characteristics of the identified population of patients with AD and comorbid insomnia was consistent with the existing literature on patients with AD and older adults with insomnia. However, it is important to note that, decreased cerebrovascular health, indicated by diagnoses of hypertension and hyperlipidemia, is itself a risk factor for developing AD [3], possibly compounding the risk of onset or degree of progression of AD among the insomnia patients.

Among older adults, insomnia severity is positively associated with anxiety, depression, and with the self-reported pain intensity and pain spreading [30]. As comorbidities, depression and insomnia overlap: about 20% of patients with insomnia also have depression [31]. For those with depression, insomnia can affect the severity of depression symptoms and the course of their mental illness [32]. For those without depression, insomnia is a risk factor of developing depression [33].

In this study, AD patients with insomnia had increased healthcare utilization compared to those without insomnia, including increased utilization of high-cost healthcare services, such as inpatient admissions, emergency room visits, and skilled nursing facility services. This included 32.8% higher adjusted odds of an inpatient admission and a 43.8% higher odds of a skilled nursing facility services claim. Prior studies of patients with AD have reported that 28%–40% of patients newly diagnosed with AD have an inpatient admission in the first year and 40%–47% have at least one emergency department visit [7, 22]; in our study the rates were comparable in magnitude with the insomnia sub-group being near the top end or slightly above those ranges. In addition, the difference in utilization between AD patients with and without insomnia was consistent with a prior study on insomnia in Medicare beneficiaries by Wickwire et al. [8]. They reported that, after adjusting for age, sex, and comorbidities (including

AD and related dementias), Medicare beneficiaries with insomnia had 61% higher odds of an inpatient admission and 49% higher odds of an ER visit than Medicare beneficiaries without insomnia [8]. Thus, the current work, combined with previous findings, indicates that the already extensive healthcare burden associated with AD patients can potentially be dramatically increased by comorbid insomnia through the increased need for high-cost services.

We found that unadjusted per-person per-year healthcare costs were \$37,356 for patients with insomnia but only \$27,990 for patients without insomnia. After matching for demographics and adjusting for, and comorbidities, this study found that the 1-year incremental cost of insomnia among patients with AD was \$10,039. We believe this may be the first such estimate in the research literature, as we could not identify any prior studies examining the incremental burden of insomnia among patients with AD. Independently, both AD and insomnia are associated with increased costs among older patients [8, 22, 34–37]. Prior estimates of the unadjusted 1-year direct healthcare costs of incident AD and related dementias have ranged from \$17,712 to \$48,985, after adjusting to 2018 USD [22, 35, 38, 39]. Estimates of the adjusted incremental cost of insomnia among older adults vary widely from \$1,143 among a Medicaid advantage population to \$63,607 among a random 5% sample of Medicare [8, 36].

Insomnia in AD patients also increases the mental and physical burden on home caregivers and reduces their overall quality of life [40–43]. This contributes to the accelerated pathway to high-cost institutionalized care observed among patients with AD and comorbid insomnia [16]. It is an important topic for future research whether treatment of insomnia can reduce the rate of cognitive decline and delaying the progression of the disease can be important in healthcare cost reduction [44, 45]. We hypothesize that effective management of insomnia in patients with AD could decrease healthcare costs by reducing the use of high-cost services such as inpatient admissions and skilled nursing facility services and reduce the burden and sleep disturbances among home caregivers. Additional research is needed to examine the relationship between effective insomnia therapy and subsequent healthcare costs.

Limitations

The limitations of this study include those inherent in any retrospective, claims-based analysis. The

MarketScan Research Databases depend on administrative claims data, and these records are subject to data coding limitations and data entry errors. Cohort misclassification based on off-label insomnia medication utilization or misdiagnosis of insomnia may have occurred. Although we used multivariable analysis to control for differences in baseline characteristics, there are several known factors, including AD severity and insomnia severity, which may influence outcomes but are not captured in administrative claims and, therefore, could not be controlled for.

In addition, non-prescription medications, such as melatonin, are not captured in claims data. As a result, we may be underestimating costs and misclassifying patients as non-insomnia patients because we are not capturing the use of melatonin as a sleep aid. Furthermore, some insomnia patients are prescribed non-hypnotic drugs (e.g., antihistamines) to improve sleep; these drugs are present in claims data but are less reliably associated with insomnia and so were not used in our patient selection process. Conversely, while the set of drugs that were used as an inclusion criterion for the insomnia cohort are all solely indicated for treating insomnia, off-label use of these medications may have resulted in the misclassification of some patients as having insomnia. As AD is often diagnosed during a hospitalization for another serious condition, all-cause costs may be overestimated for patients whose insomnia diagnosis occurred on or near their initial AD diagnosis [46]. While this may inflate the overall costs estimated in this study, both cohorts should be impacted similarly due to the patient matching and the method of assigning a proxy insomnia diagnosis date for the control cohort.

Finally, as this study was limited only to those individuals with commercial health coverage or private Medicare supplemental coverage. The results of this analysis may not be generalizable to AD or insomnia patients with other insurance or without health insurance coverage.

CONCLUSIONS

Comorbid insomnia is associated with a significant healthcare burden among patients with AD. In this claims-based study, patients with AD and comorbid insomnia had greater utilization of high-cost services such as inpatient admissions and skilled nursing facility services than patients without comorbid insomnia. This corresponded with higher total

healthcare costs among AD patients with comorbid insomnia. These were true across a wide range of medical setting categories including inpatient admissions, emergency room visits, outpatient office visits, and costs associated with skilled nursing facility care. Prior research has examined AD and insomnia separately, but this research helps to characterize the population of AD patients with comorbid insomnia. The burden of AD alone is substantial. These findings suggest that comorbid insomnia only creates additional burden to the healthcare system, in addition to increasing caregiver burdens at home and increasing need for institutionalization. There is clear unmet need for properly managing insomnia in patients with underlying AD or dementia. Slowing the progression of the underlying neurological factors of dementia and treating symptoms that cause additional issues, such as insomnia, should be the focus of new treatment options. Understanding this population, including the interactions of insomnia with AD and associated treatment needs, has potential to improve the progression of the disease and delay or reduce the need for institutional care.

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DATA AVAILABILITY STATEMENT

The data that support the findings of this study are available from IBM Watson Health. Restrictions apply to the availability of these data, which were used under license for this study.

SUPPLEMENTARY MATERIAL

The supplementary material is available in the electronic version of this article: <https://dx.doi.org/10.3233/JAD-210713>.

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